# **RESEARCH ARTICLE**

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# "Childhood Anemia in India: an application of a Bayesian geo-additive model"



Holendro Singh Chungkham<sup>1,2</sup>, Strong P. Marbaniang<sup>3,4\*</sup> and Pralip Kumar Narzary<sup>5</sup>

#### **Abstract**

**Background:** The geographical differences that cause anaemia can be partially explained by the variability in environmental factors, particularly nutrition and infections. The studies failed to explain the non-linear effect of the continuous covariates on childhood anaemia. The present paper aims to investigate the risk factors of childhood anaemia in India with focus on geographical spatial effect.

**Methods:** Geo-additive logistic regression models were fitted to the data to understand fixed as well as spatial effects of childhood anaemia. Logistic regression was fitted for the categorical variable with outcomes (anaemia (Hb < 11)) and no anaemia  $(Hb \ge 11)$ ). Continuous covariates were modelled by the penalized spline and spatial effects were smoothed by the two-dimensional spline.

**Results:** At 95% posterior credible interval, the influence of unobserved factors on childhood anaemia is very strong in the Northern and Central part of India. However, most of the states in North Eastern part of India showed negative spatial effects. A U-shape non-linear relationship was observed between childhood anaemia and mother's age. This indicates that mothers of young and old ages are more likely to have anaemic children; in particular mothers aged 15 years to about 25 years. Then the risk of childhood anaemia starts declining after the age of 25 years and it continues till the age of around 37 years, thereafter again starts increasing. Further, the non-linear effects of duration of breast-feeding on childhood anaemia show that the risk of childhood anaemia decreases till 29 months thereafter increases.

**Conclusion:** Strong evidence of residual spatial effect to childhood anaemia in India is observed. Government child health programme should gear up in treating childhood anaemia by focusing on known measurable factors such as mother's education, mother's anaemia status, family wealth status, child health (fever), stunting, underweight, and wasting which have been found to be significant in this study. Attention should also be given to effects of unknown or unmeasured factors to childhood anaemia at the community level. Special attention to unmeasurable factors should be focused in the states of central and northern India which have shown significant positive spatial effects.

Keywords: Spatial effects, Geo-additive logistic regression, P-splines, Childhood anaemia

#### **Background**

Anemia among children is still a major public health concern in both developed and developing countries. Anemia is a condition in which the number and size of red blood cells or haemoglobin concentration is lower than the established cut-off value [1]. Haemoglobin is

essential to carry oxygen and if the body has abnormal or low red blood cells or not enough haemoglobin level, there will be a reduced capacity of the blood to carry oxygen to the body tissues. Globally, anemia affects 1.6 billion people, of which 47.4% were preschool-age children [2]. According to the World Health Organization, (2008), anemia is considered a severe public health problem if the prevalence is 40% or more [2]. In India, 58.5% percent of children between the age of 6 months to 5 years were anemic during 2015–2016 [3]. Moreover, studies

<sup>&</sup>lt;sup>4</sup> Department of Statistics, Sankardev College, Shillong, Meghalaya, India Full list of author information is available at the end of the article



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<sup>\*</sup>Correspondence: marbaniangstrong@gmail.com

have acknowledged high prevalence of anaemia in low and middle-income countries [4], with 67.6 and 65.6% preschool-age children in Africa and South-East Asia suffered from anaemia [2] respectively.

Iron is an essential element of haemoglobin, and iron deficiency is the most common cause of anaemia. However, deficiency in micronutrient-rich diet, Vitamin A, and Vitamin B12 could be the reason for iron deficiency [5]. Also, diseases like diarrhea [6], malaria [7], helminth infection, and hookworms [5] increases the risk of anemia. In India, due to various socio-economic, cultural, and religious beliefs, dietary food habits also vary across the population. Dietary pattern is an essential factor associated with iron intake and absorption. For example, a vegetarian diet may increase the risk of anemia due to lack of iron fortification [8]. Existing literature have also shown that socio-economic factors such as lower maternal education, low economic status [9], and demographic factors such as age and sex of a child [10] affect anaemia. Maternal health status during pregnancy had a significant impact on the health and nutritional status of the child. Evidence from previous studies reported that maternal anaemia, and child nutritional statuses such as wasting, stunting and underweight increase the risk of anaemia [11, 12]. During the first 5 years of life, children are most vulnerable to iron-deficiency anaemia because of increased iron requirements due to their rapid growth [13]. Iron deficiency in children is a serious concern because it may increase childhood morbidity, impaired growth development, and have long term effects on cognitive development and school performance [13].

Accounting for geographical heterogeneity of anaemia and the possible cause of heterogeneity is vital for the allocation of health resources to prevent and control anaemia. Geographical heterogeneity can be an effect of an unobserved independent variables which may include contextual factors. According to Koissi & Högnäs, (2013) ignorance of geographical heterogeneity due to unobserved characteristics could lead to biased estimation of parameters [14]. Geographical heterogeneity could be the effect of the unmeasured factors, which means that the geographical differences of factors that caused anaemia can be partially explained by the variability in environmental factors [15]. Environmental factors such as availability of toilet facility, type of house, source of drinking water, seasonality influence the risk of anaemia among children. Studies found that lower odds of anemia among children living in household with better toilet facility, improve drinking water and better housing condition [16]. Malaria which causes anaemia is known to be associated with altitude and weather conditions such as temperature and rainfall [17]. Similarly, soil-transmitted helminth infection, which causes anaemia is influenced by the distance to water bodies, surface temperature, index of vegetation and rainfall [18]. There are number of studies using different statistical models such as multilevel and spatial mixed model to determine the effect of geographical heterogeneity on childhood anaemia in India [9, 10]; however, all these studies have overlooked the advantage of using bivariate spline in modelling geographical heterogeneity. Above models failed to explain especially the non-linear effect of continuous covariates on childhood anaemia. Thus, the pioneering contribution of this study would be to explore correlated spatial effect of anaemia among children aged 6 to 59 months using the spatial mixed model by assuming the flexible approach of bivariate splines. This study would probably be the first in India to map childhood anaemia in terms of residual spatial effects due to unmeasured factors. So, the map would have important implications for targeted policy for allocation of resources and to search for unmeasured variables that are responsible for residual spatial effects.

# **Methods**

### Study area

The study used the fourth round of the Indian National Family and Health Survey (2015–2016) which adopted a multi-stage stratified cluster sampling design [19]. From all over India, total of 699,686 eligible women between 15 and 49 years of age completed the interview. The data for the present study uses child as the unit of analysis, rather than the mother. Information was available on 259,627 children born in the last 5 years preceding the survey. The present study excluded the two union territories i.e., Andaman & Nicobar Island and Lakswdeep as their borders are not connected to other parts of India and which would create problem in the estimation of spatial effects. Children with missing haemoglobin level were also dropped from the analysis. With this criterion the final analytical sample size consists of children 208,707.

## Outcome variable and covariates

The outcome variable used in the analysis was based on the categorization of haemoglobin level of children adjusted for altitude. The children whose haemoglobin level was less than 11Hb categorised as being anaemic otherwise not anaemic. The covariates in the present study were selected based on previous study [15] and theoretical understanding of the issue under investigation. As such, mother's educational level, anaemia status, age and duration of breast feeding are considered as covariates. Children related characteristics considered are whether children had cough, had fever, received vitamin A, whether stunting, wasting, underweight, birth weight, birth order, and age of the children. Further, household wealth index and family size are included in the study.

Duration of breast feeding, age of children, and mother's age were treated as continuous variables. However, the standard -2SD cut off values of z-scores categorization of height for age, weight for height, and weight for age were used to characterize stunting, wasting and underweight respectively.

#### Statistical analysis

Multiple logistic regression model was employed to select potential covariates for childhood anaemia prior to spatial analysis. A significance level of 20% was set for the selection of potential covariates to allow for selection of more variables to be used in the further analysis of spatial modelling. Geo-additive logistic regression models were fitted to the data to understand fixed as well as spatial effects of childhood anaemia. Basically, the model takes the form of a multiple variable hierarchical model as

$$g(\mu) = U\beta + e_i$$

Where g is the logit link function which gives the log odds of being anaemic and it links the mean of the response to the predictor  $U\beta + e_i$ , and  $e_i$  is the area level random effects representing unmeasured contextual factors. More formally, we can formulate the above model as, if  $p_{ij}$  is the probability that child j from location i being anaemic, then child anaemic status which is binary is distributed as  $Bernoulli(p_{ij})$ . Then, following models were fitted to estimate fixed and spatial effects.

$$M0: logit(p_{ij}) = z_i'\beta$$

$$M1: logit(p_{ij}) = z'_i\beta + f_1(u_{i1}) + f_1(u_{i2}) + \dots + f_1(u_{ip})$$

$$M2: logit(p_{ij}) = z'_i \beta + f_{spatial}(S_i)$$

$$M3: logit(p_{ij}) = z_i'\beta + f_1(u_{i1}) + f_1(u_{i2}) + \dots + f_1(u_{ip}) + f_{spatial}(S_i)$$

All categorical and continuous variables were treated as fixed effects in M0. In case of M1, categorical variables were employed as fixed effects and continuous variables were modelled by non-parametric smooth functions  $f_js$ . Model M2 included a spatial effect of the state where a child belongs in addition to the fixed effects of categorical variables. Finally, M3 was a combination of M1 and M2. The smooth functions  $f_js$  were specified as Bayesian splines and can be approximated by polynomial spline priors of degree l at equally spaced knots  $u_j^{min} = \gamma_{j0}, \gamma_{j1}, \ldots, \gamma_{js} = u_j^{max}$  which are within the domain of covariate  $u_j$ , and the spatial component  $f_{spatial}(S_i)$  with Markov random field prior [20, 21] which captures the area of the child random effect. The

Bayesian spline can be expressed as a linear combination of d = s + l basis functions  $B_m$  having the form as,

$$f_j(u_j) = \sum_{m=1}^d \varepsilon_{jm} B_m(u_j)$$

Then, the Bayesian estimation of the above spline reduces to estimating model parameters  $\varepsilon_j s$  by assigning first or second order random walk priors for the regression coefficients. A tensor product of two-dimensional spline has been used to model the spatial effect as,

$$f_{spatial}(u_1, u_2) = \sum_{i}^{k} \sum_{j}^{k} B_{spatial,ij} B_{1i}(u_1) B_{2j}(u_2)$$

where, the combination  $(u_1,u_2)$  corresponds to the coordinates of the location of the data point, latitude and longitude, or the location centroids based on the map. The commonly available spatial smoothness priors in spatial statistics [22] based on the four nearest neighbours have been adopted.

A fully integrated Bayesian approach was adopted to estimate the parameters and the estimated posterior odds ratio (OR) can be interpreted as the odds ratio from the logistic regression models. The models were fitted using the freely available package bamlss [23] in R (R Core Team, 2020). A total of 40,000 MCMC iterations and 10,000 number of burn in samples were used in the analysis. Convergence of models were checked through autocorrelations and sampling paths. Finally, models were compared by Deviance Information Criterion (DIC) values [24], where the model with the smallest value is the preferred one. The DIC is calculated as  $DIC = \overline{D} + p_D$ , where  $\overline{D}$  is the posterior mean of the model deviance, which gives a measure of goodness of fit, and  $p_D$  is the effective number of parameters describing the complexity of the model and controls for penalty for model overfitting.

# Results

# **Descriptive results**

Table 1 provides prevalence of childhood anaemia according to region and states in India. Northern, central, and eastern regions show high prevalence of anaemia compared to other regions. The prevalence is above 60% in these three regions. The states of Chandigarh and Haryana show relatively high prevalence of anaemia of about 73 and 72% in northern region. In the central region, Madhya Pradesh and Uttar Pradesh show relatively high prevalence of anaemia. Jharkhand and Bihar are the states in eastern region having relatively high prevalence of anaemia of about 70 and 64% respectively. Most of the

**Table 1** State variation of childhood anaemia

| Region/State         | Percentage of<br>children<br>(Anaemic) | Number of<br>cases with<br>anaemic<br>children |
|----------------------|--|--|
| Northern             | 62.2                                   | 21,765   |
| Chandigarh           | 72.7                                   | 112  |
| Haryana              | 72.3                                   | 4725   |
| Himachal Pradesh     | 58.1                                   | 1324   |
| Jammu and Kashmir    | 59.6                                   | 3986   |
| Delhi                | 61.3                                   | 627  |
| Punjab               | 57.3                                   | 2544   |
| Rajasthan            | 60.9                                   | 8447   |
| Central              | 63.0                                   | 41,351   |
| Chhattisgarh         | 42.9                                   | 3060   |
| Madhya Pradesh       | 69.7                                   | 14,015   |
| Uttar Pradesh        | 63.8                                   | 21,468   |
| Uttarakhand          | 59.1                                   | 2808   |
| Eastern              | 61.2                                   | 27,158   |
| Bihar                | 63.6                                   | 13,332   |
| Jharkhand            | 70.1                                   | 7002   |
| Odisha               | 48.6                                   | 4393   |
| West Bengal          | 55.6                                   | 2431   |
| North-Eastern        | 35.8                                   | 10,504   |
| Arunachal Pradesh    | 53.3                                   | 1956   |
| Assam                | 35.7                                   | 2838   |
| Manipur              | 24.2                                   | 1153   |
| Meghalaya            | 48.7                                   | 1706   |
| Mizoram              | 23.9                                   | 975  |
| Nagaland             | 26.2                                   | 908  |
| Sikkim               | 56.6                                   | 457  |
| Tripura              | 48.1                                   | 511  |
| Southern             | 54.7                                   | 10,806   |
| Andhra Pradesh       | 58.2                                   | 1246   |
| Karnataka            | 62.1                                   | 3818   |
| Kerala               | 36.0                                   | 742  |
| Puducherry           | 43.8                                   | 408  |
| Tamil Nadu           | 51.5                                   | 3461   |
| Telangana            | 64.3                                   | 1131   |
| Western              | 58.2                                   | 8524   |
| Dadra & Nagar Haveli | 83.9                                   | 220  |
| Daman & Diu          | 72.4                                   | 205  |
| Goa                  | 48.3                                   | 174  |
| Gujarat              | 63.7                                   | 3839   |
| Maharashtra          | 52.9                                   | 4086   |
| India                | 57.6                                   | 120,108  |

states in the north-eastern region show comparatively low prevalence of anaemia ranging from 24 to 57%. The states of Karnataka and Telangana show relatively high prevalence of anaemia above 60%. The overall prevalence of anaemia in India is about 58%.

Table 2 provides a comparison of childhood anaemia across categorical covariates and a test of significance difference between categories of each covariate by chisquare test. It is evident that children from rural, mother with low education, household of poor economic condition show higher prevalence of anaemia than their respective counterparts. There is a clear significant difference in childhood anaemia by place of residence, mother's education and household wealth. But no significant difference in childhood anaemia by sex of child is observed. Children with fever show a tendency of higher prevalence of anaemia. It can also be seen that consumption of vitamin A supplement during childhood is helpful to reduce prevalence of anaemia. Under nutrition of children also show an increase in prevalence of anaemia. At 5% level of significance the categorical variables- place of residence, mother's education, mother's anaemic status, household economic status, children's fever, vitamin A, stunting, wasting, and underweight are associated with childhood anaemia without controlling for other covariates. The categorical variables children's birth order, children's birth weight and household size show a non-significant effect on childhood anaemia at 20% level of significance in the preliminary analysis. Therefore, only categorical variables listed in Table 2 are included in the spatial logistic regression model in Table 4.

#### Model selection

The selection of the most preferred model is based on the deviance information criterion (DIC) and deviance values. Model with the smallest values of DIC and deviance is the preferred model. With these criteria, model M3 is the preferred model (Table 3). Therefore, interpretations of results (Table 4) and discussions are based on model M3.

## **Fixed effects**

Table 4 shows fixed effects to childhood anaemia. Place of residence, mother's education, poorest, rich, richest categories of household wealth, fever, cough, child under nutrition and mother's anaemic status are fixed effects variables which are significant to childhood anaemia. The fixed effects coefficient for fever is positive, which indicates that children with fever are likely to increase the risk of childhood anaemia. Children who take vitamin A supplement decrease the likelihood of becoming anaemic. Children from rich or richest quintile of household wealth also have lesser risk of childhood anaemia than those who belong to poorest quintile. Children who are malnourished increase the risk of childhood anaemia. Mother's anaemic status has a positive effect on childhood anaemia. This means

**Table 2** Prevalence of childhood anaemia by fixed covariates with effect coding used in model

Factor N (%) Р\* **Effect coding** Place of residence < 0.001 Urban 27,338 (55.2) 1 -1<sup>R</sup> Rural 92,770 (58.3) Sex of the child 0.644 62,486 (57.5) Male -1<sup>R</sup> Female 57,622 (57.6) < 0.001 Mother's education Primary 17,845 (58.3) 1 Secondary 50,460 (54.1) 2 Higher 9467 (50.1) 3 No education -1<sup>R</sup> 42,336 (64.3) Wealth index < 0.001 Poor 28,395 (57.6) 1 Middle 23,422 (56.2) 2 Rich 18,677 (53.9) 3 Richest 14,804 (52.9) 4 -1<sup>R</sup> Poorest 34,810 (63.2) < 0.001 Fever Yes 16,729 (60.9) -1<sup>R</sup> 103,295 (57.1) No 84 (52.8) Missing Cough 0.220 13,887 (57.1) Yes -1<sup>R</sup> No 106,159 (57.6) Missing 62 (54.9) Child received vitamin A > 0.001 Yes 38,674 (58.1) -1<sup>R</sup> No 80,003 (57.3) Missing 1431 (57.1) Stunting < 0.001 Yes 50,438 (62.7) -1<sup>R</sup> No 64,015 (53.6) 5655 (63.9) Missing Underweight < 0.001 Yes 45,252 (63.7) -1<sup>R</sup> No 69,201 (53.7) Missing 5655 (63.9) Wasting < 0.001 Yes 8814 (64.1) -1<sup>R</sup> No 105,639 (56.8) Missing 5655 (63.9) < 0.001 Mother anaemic Yes 48,928 (67.8) -1<sup>R</sup> No 70,787 (52.1) Missing 393 (58.1)

**Table 3** Model comparison by deviance information criterion (DIC)

| Model Fit | Deviance   | pD    | DIC        |
|-----------|------------|-------|------------|
| M0        | 171,173.90 | 19.79 | 171,154.10 |
| M1        | 170,885.30 | 37.71 | 170,847.60 |
| M2        | 165,233.90 | 51.77 | 165,182.10 |
| M3        | 164,909.50 | 69.92 | 164,839.60 |

Table 4 Fixed effects on childhood anaemia in India

| Variable                               | Mean     | SD    | 10%      | Median   | 90%             |
|--|----------|-------|----------|----------|-----------------|
| Place of residence                     |          |       |          |          |                 |
| Rural <sup>R</sup>                     |          |       |          |          |                 |
| Urban                                  | 0.0359*  | 0.008 | 0.0262   | 0.0355   | 0.0461          |
| Sex of child                           |          |       |          |          |                 |
| Female <sup>R</sup>                    |          |       |          |          |                 |
| Male                                   | 0.0074   | 0.006 | -0.0003  | 0.0075   | 0.0148          |
| Mother's education                     | n        |       |          |          |                 |
| No education <sup>R</sup>              |          |       |          |          |                 |
| Primary                                | 0.0563*  | 0.014 | 0.0386   | 0.0564   | 0.0740          |
| Secondary                              | -0.0358* | 0.010 | -0.0481  | -0.0361  | - 0.0229        |
| Higher                                 | -0.1843* | 0.016 | - 0.2056 | -0.1844  | <b>-</b> 0.1625 |
| Wealth index                           |          |       |          |          |                 |
| Poorest <sup>R</sup>                   |          |       |          |          |                 |
| Poor                                   | 0.0740*  | 0.012 | 0.0585   | 0.0736   | 0.0893          |
| Middle                                 | 0.0069   | 0.012 | -0.0079  | 0.0071   | 0.0222          |
| Rich                                   | -0.0904* | 0.013 | -0.1072  | - 0.0904 | - 0.0736        |
| Richest                                | -0.1332* | 0.017 | -0.1548  | -0.1330  | <b>-</b> 0.1125 |
| Child had fever<br>No <sup>R</sup>     |          |       |          |          |                 |
| Yes                                    | 0.0326*  | 0.010 | 0.0200   | 0.0327   | 0.0451          |
| Child had cough<br>No <sup>R</sup>     |          |       |          |          |                 |
| Yes                                    | -0.0594* | 0.010 | -0.0723  | - 0.0596 | - 0.0466        |
| Child received vita<br>No <sup>R</sup> | min A    |       |          |          |                 |
| Yes                                    | -0.0041  | 0.007 | -0.0125  | - 0.0042 | 0.0042          |
| Child stunted<br>No <sup>R</sup>       |          |       |          |          |                 |
| Yes                                    | 0.0999*  | 0.007 | 0.0903   | 0.1000   | 0.1091          |
| Child underweight                      |          | 0.007 | 0.0903   | 0.1000   | 0.1051          |
| Yes                                    | 0.0797*  | 0.008 | 0.0698   | 0.0795   | 0.0899          |
| Child wasted                           | 0.07 57  | 0.000 | 0.0070   | 0.07 73  | 0.0077          |
| No <sup>R</sup>                        |          |       |          |          |                 |
| Yes                                    | 0.0387*  | 0.012 | 0.0235   | 0.0387   | 0.0541          |
| Mother anaemic                         | 3.0307   | 5.012 | 0.0233   | 0.0307   | 5.05 11         |
| INO                                    |          |       |          |          |                 |

 $<sup>^{\</sup>rm R}$  : Reference category. \*:Statistically significant at 5% alpha

 $<sup>^{\</sup>rm R}$  : Reference category; \*: p-value of chi-square test of independence

that children whose mothers are anaemic have higher risk of being anaemic than those whose mothers are not anaemic.

#### Non-linear effects

Another reason behind the geo-additive modelling is the ability to incorporate non-linear effects of continuous variables in the model. In the present study, we incorporated non-linear effects of age of child, mother's age and, duration of breast feeding.

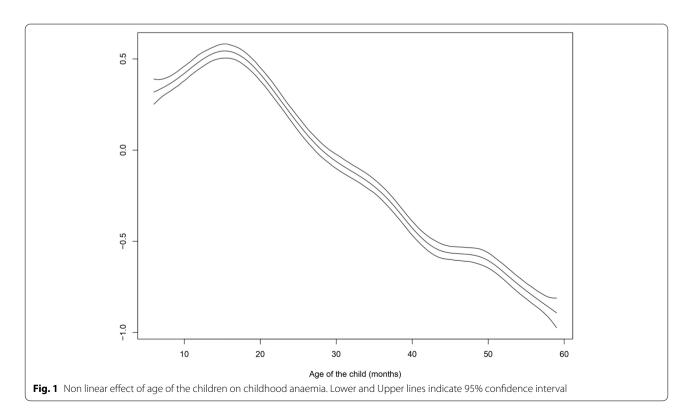
The age of children has non-linear effect on childhood anaemia (Fig. 1). It is evident from Fig. 1 that as the age of children increases, its effect on childhood anaemia decreases, which indicates, older children are less likely to have the risk of childhood anaemia. The risk of having anaemia is much higher among younger children aged about 6 months to about 15 months and decreases thereafter.

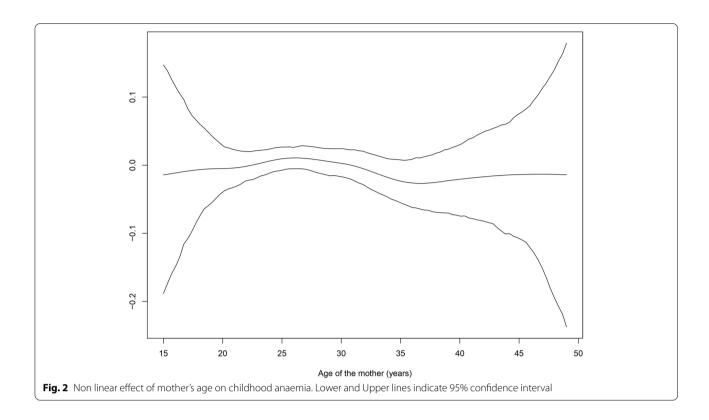
Mother's age also has a non-linear effect on child-hood anaemia (Fig. 2). The functional relationship between childhood anaemia and mother's age depicts almost a U shape pattern. This indicates that mothers of young (in particular mothers aged 15 years to about 25 years) and old ages are more likely to have children who are anaemic. The risk of childhood anaemia starts declining after the age of 25 years and continuous till the age of around 37 years, thereafter again starts increasing.

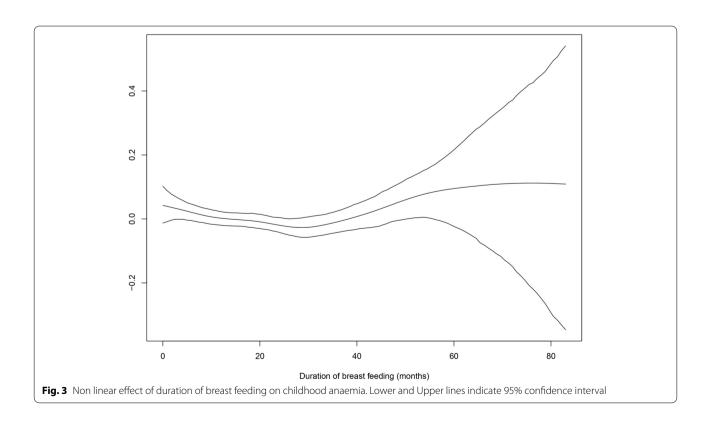
Figure 3 shows the non-linear effects of duration of breast feeding on childhood anaemia. The risk of childhood anaemia decreases till 29 months, thereafter increases. This indicates improvement in childhood anaemia with increase in duration of breast feeding. The credible intervals are wider at extreme ages because of small cases of observations.

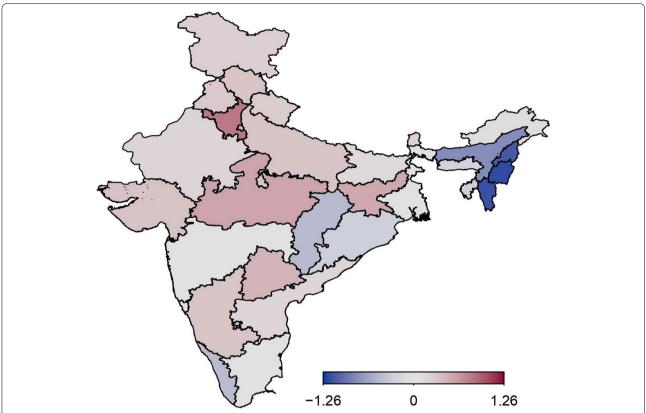
## **Spatial effects**

Figure 4 displays the estimates of the spatial effects of childhood anaemia, with colour range goes from blue to red representing low to high risk of childhood anaemia. Spatial effects represent unobserved influences, such as environmental and climatic factors, availability of good transport facility, and access to good services for child health. The figure clearly shows evidence of residual spatial effects of childhood anaemia in India with most of states showing significant positive/negative effects with respect to the 95% posterior credible interval map (Fig. 5). With respect to 80% posterior credible interval more states show significant spatial effects (Fig. 6). Most of the states in northern and central regions show significant positive spatial effects with respect to 95% credible interval. However, almost all states in north-eastern region of India show significant negative spatial effects with regard to the 80% credible interval (Fig. 6).









**Fig. 4** Residual spatial effect to childhood anaemia. Colour ranges from blue to red representing low to high risk of childhood anaemia. Source of Shapefile Map: Bhuvan India Geo Platform of Indian Space Research Organisation, Govt. of India

#### Discussion

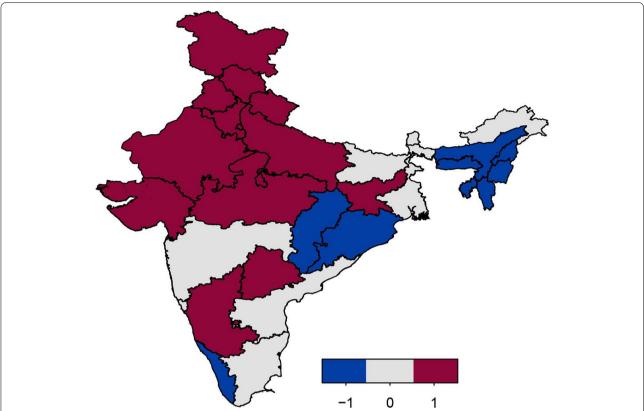
In India Childhood anaemia cuts across all the sections of society with varying intensity. Its prevalence, as per the WHO classification, is a severe public health problem for India. Except for Mizoram, Manipur, Nagaland, Assam, and Kerala for all the states and union territories (UTs,) anaemia is a matter of concern, whereas for states like Haryana, Jharkhand, and Madhya Pradesh it is of extremely serious concern. These three states need to revisit existing programs targeting to address the child health in general and anaemia in particular.

Anaemia has a close link with the food habit. Food habit is closely associated with culture and the nature. Geographical settings decide the nature of food supply and the micronutrients. Within the same geographical settings culture may encourage or discourage some group of population to consume or avoid certain nutritious food. For example tribal culture of northeast India approves consumption of varieties of insects, whereas for non-tribals consumption of such insects is considered as taboo. Probably because of this reason the tribal dominated states like Mizoram, Manipur,

and Nagaland have very low prevalence of anaemic children. However, our finding contradicts other studies in India that children from lowest socioeconomic strata have more likelihood of suffering from anaemia [9, 25] and Nepal [26].

The prevalence of anaemia among children in rural areas is comparatively higher than their counter part in India. Rural mass in India might be less aware about the balanced diet which has potentials to improve the hemoglobin count. Because, as high as one third of rural population in India are illiterate. Ignorance of food items relating to iron content food staff may also add to the problem of anaemia in rural areas. This indicates that mass media campaign to address anaemia should emphasize on pictorial depiction and or audiovisual means, rather than on the written leaflets. A distinct negative relationship between wealth quintile and child anaemia is quite evident. This is indicative of the fact that economically poorer households may not be able to afford to procure food regularly and especially the nutritious food times. This calls for better Public Distribution System (PDS) which provides subsidized

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**Fig. 5** The 95% credible intervals map for prevalence of anaemia. Blue: negative effect; light gray: insignificant effect; red: positive effect. Source of Shapefile Map: Bhuvan India Geo Platform of Indian Space Research Organisation, Govt. of India

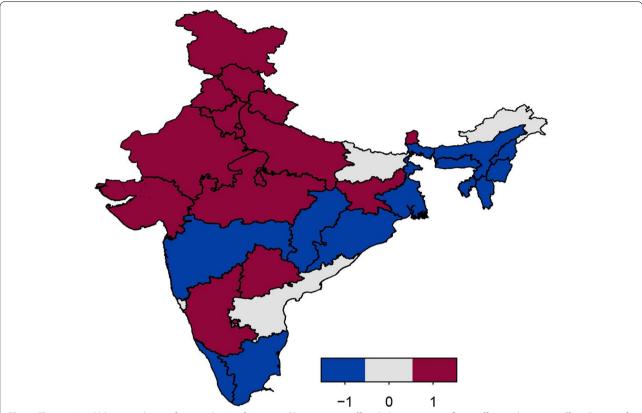
food in India. The system needs to keep an eye on mainly on regularity, quantity, and quality of supplies.

Uneducated mothers are less equipped with knowledge of hygiene and proper knowledge of child care. Unhealthy feeding habit can lead to various types of food related health problems. Feeding practice is closely associated with diarrhoeal disease and studies exhibit that there is positive relationship between diarrhoea and anemia. Unlike earlier studies [8, 10] no significant association is noted between sex of the child and prevalence of anaemia in the present study. Children who take vitamin A supplement decrease the likelihood of becoming anaemic. But earlier study [8] did not find significant statistical association between vitamin A intake and childhood anaemia. In India, poor and illiterate families leave their baby on the mud floor. The crawling baby in absence of a care taker may put to mouth anything it comes to her/his hand. Such activities may lead to various infections and morbidities, for which younger children have more likelihood of suffering from anaemia. Other studies also indicate that younger children have more chances of having anaemia [15, 26]. Very young mothers definitely are less educated and relatively old mothers might take child rearing for granted, as they may already have older children and experienced of child rearing. Other study also indicates U-shape relationship between mother's age and the childhood anaemia [15] and others [10, 27] found children born to young mothers are more likely to be anaemic. In India usually the educated and rich women, due to various reasons, do not practice exclusive breast feeding. Exclusive breast feeding in India is usually practiced among the less educated and poor women, as a result a positive association between exclusive breast feeding and childhood anaemia is observed. However, this finding contradicts studies conducted elsewhere [28].

## Limitations

The present study is not without any limitation despite using an innovative statistical technique. First, our study is based on cross-sectional design. Therefore, control of major confounders and no causal inferences can be made in spite of robustness in the analysis. Second, the study uses only relevant variables in our data set leading to

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**Fig. 6** The 80% credible intervals map for prevalence of anaemia. Blue: negative effect; light gray: insignificant effect; red: positive effect. Source of Shapefile Map: Bhuvan India Geo Platform of Indian Space Research Organisation, Govt. of India

omission of certain important variables such as clamping of umbilical cord after birth mentioned in some studies.

## **Conclusions**

There is strong evidence of residual spatial effect to childhood anaemia in India. Government child health programme should gear up in treating childhood anaemia by focusing on known measurable factors such as mother's education, mother's anaemia status, family wealth status, child fever, stunting, underweight, and wasting which have been found to be significant in this study. Attention should also be given to effects of unknown or unmeasured factors of childhood anaemia at the community level. Special attention to these unmeasurable factors should be focused in the states of central and northern India which have shown significant positive spatial effects. As the problem of anemia is multi-faceted, the Anemia Mukt Bharat strategy adopted under Poshan Abhiyaan shows great hope in bringing down the prevalence of anemia in India by adopting 6x6x6 strategy [29]. The strategy of targeting six groups of population, six interventions, and six institutional mechanisms is very fascinating but only time will tell its success.

## Abbreviation

DIC: Deviance information criterion; WHO: World health organisation; OR: Odds ratio; UTs: Union territories; HSC: Holendro singh chungkham; SPM: Strong P marbaniang; PKN: Pralip kumar narzary.

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## Authors' contributions

HSC, SPM conceived the study, involved in the study design, data analysis, interpret the data, drafted the manuscript. PKN drafted and critically reviewed the manuscript. All authors read and agreed on the submitted final manuscript. The author(s) read and approved the final manuscript.

#### Author's information

HSC is an Associate Professor at the Indian Statistical Institute (ISI), North-East Centre, Tezpur, Assam, India and he is also a Researcher at Stress Research Institute, Stockholm University. SPM is a Research Scholar in the Department of Public Health and Mortality Studies, International Institute for Population Sciences, Mumbai, Maharashtra, India and he is also a faculty in the Department of Statistics, Sankardev College, Shillong, Meghalaya, India. PKN is an Associate Professor in the Department of Geography, Bodoland University, Kokrajhar, Assam, India.

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#### Availability of data and materials

The datasets generated and/or analysed during the current study are available in the Website of Demographic Health Survey https://dhsprogram.com/methodology/survey/survey-display-355.cfm. We submitted a request to the DHS by mentioning the objectives of this study and thereafter was granted the permission to download the dataset.

## **Declarations**

#### Ethics approval and consent to participate

The 2015–16 Indian Demographic Health Survey data are available to the public by request from the DHS website https://dhsprogram.com/methodology/survey/survey-display-355.cfm. The available is without identifier. As such there is no need for ethical approval and consent to participate does not arise.

#### Consent for publication

Not applicable.

#### Competing interests

The authors declare that they have no competing interests.

#### **Author details**

<sup>1</sup>Indian Statistical Institute, North-East Centre, Tezpur, Assam, India. <sup>2</sup>Stress Research Institute, Stockholm University, Stockholm, Sweden. <sup>3</sup>Department of Public Health & Mortality studies, International Institute for Population Sciences, Mumbai, Maharashtra, India. <sup>4</sup>Department of Statistics, Sankardev College, Shillong, Meghalaya, India. <sup>5</sup>Department of Geography, Bodoland University, Kokraihar, Assam, India.

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## References

- Kotwal A. Iron deficiency anaemia among children in South East Asia: determinants, importance, prevention and control strategies. Curr Med Res Pract. 2016;6(3):117–22 Available from: http://linkinghub.elsevier. com/retrieve/pii/S2352081716300654.
- Benoist B de, McLean E, Egli I, Cogswell M, editors. Worldwide prevalence of anaemia 1993-2005: WHO global database on anaemia. World Health Organization; 2008 [cited 2020 Aug 10]. Available from: https://www.who. int/vmnis/anaemia/prevalence/en/
- International institute for population sciences (IIPS) and ICF. National Family Health Survey (NFHS-4), 2015–16: India Mumbai:IIPS; 2017. Available from: http://rchiips.org/nfhs/NFHS-4Reports/India.pdf
- McLean E, Cogswell M, Egli I, Wojdyla D, de Benoist B. Worldwide prevalence of anaemia, WHO vitamin and mineral nutrition information system, 1993–2005. Public Health Nutr. 2009;12(04):444 Available from: http://www.journals.cambridge.org/abstract\_S1368980008002401.
- Balarajan Y, Ramakrishnan U, Özaltin E, Shankar AH, Subramanian
  Anaemia in low-income and middle-income countries. Lancet.
  2011;378(9809):2123–35 Available from: https://linkinghub.elsevier.com/retrieve/pii/S0140673610623045.
- Howard CT, de Pee S, Sari M, Bloem MW, Semba RD. Association of diarrhea with anemia among children under age five living in rural areas of Indonesia. J Trop Pediatr. 2007;53(4):238–44 Available from: https://academic.oup.com/tropej/article-lookup/doi/10.1093/tropej/fmm011.
- Calis JCJ, Phiri KS, Faragher EB, Brabin BJ, Bates I, Cuevas LE, et al. Severe Anemia in Malawian children. N Engl J Med. 2008;358(9):888–99 Available from: http://www.nejm.org/doi/abs/10.1056/NEJMoa072727.
- Ray R. Mother's autonomy and child anemia: A case study from India. Child Youth Serv Rev. 2020;112(March):104537. Available from: https://doi.org/https://doi.org/10.1016/j.childyouth.2019.104537.
- Sharma H, Singh SK, Srivastava S. Socio-economic inequality and spatial heterogeneity in anaemia among children in India: Evidence from NFHS-4 (2015–16). Clin Epidemiol Glob Heal. 2020;8(4):1158–1171. Available from: https://doi.org/https://doi.org/10.1016/j.cegh.2020.04.009.
- Dutta M, Bhise M, Prashad L, Chaurasia H, Debnath P. Prevalence and risk factors of anemia among children 6–59 months in India: A multilevel

- analysis. Clin Epidemiol Glob Heal. 2020;8(3):868–878. Available from: https://doi.org/https://doi.org/10.1016/j.cegh.2020.02.015.
- Rahman MS, Mushfiquee M, Masud MS, Howlader T. Association between malnutrition and anemia in under-five children and women of reproductive age: Evidence from Bangladesh Demographic and Health Survey 2011. Adu-Afarwuah S, editor. PLoS One. 2019;14(7):e0219170. Available from: https://dx.plos.org/https://doi.org/10.1371/journal.pone.0219170
- Engidaye G, Melku M, Yalew A, Getaneh Z, Asrie F, Enawgaw B. Under nutrition, maternal anemia and household food insecurity are risk factors of anemia among preschool aged children in Menz Gera Midir district, eastern Amhara, Ethiopia: a community based cross-sectional study. BMC Public Health 2019;19(1):968. Available from: https:// bmcpublichealth.biomedcentral.com/articles/https://doi.org/10.1186/ s12889-019-7293-0
- 13. WHO. Guideline: Intermittent iron supplementation in preschool and school-age children. Geneva: World Health Organization, Geneva; 2011. Available from: https://www.who.int/nutrition/publications/micronutrients/guidelines/guideline\_iron\_supplementation\_children/en/
- Koissi M-C, Högnäs G. Using WinBUGS to study family frailty in Child mortality, with an application to Child survival in Ivory Coast. African Popul Stud. 2013;20(1):1–17 Available from: http://aps.journals.ac.za/pub/article/ view/384.
- Ngwira A, Kazembe LN. Bayesian random effects modelling with application to childhood anaemia in Malawi. BMC Public Health 2015 ;15(1):161. Available from: http://bmcpublichealth.biomedcentral.com/ articles/https://doi.org/10.1186/s12889-015-1494-y
- Baranwal A, Baranwal A, Roy N. Association of Household Environment and Prevalence of Anemia among children Under-5 in India. Front Public Heal 2014 20;2. Available from: http://journal.frontiersin.org/ article/https://doi.org/10.3389/fpubh.2014.00196/abstract.
- Marbaniang SP, Ladusingh L. Meteorological conditions and malaria cases—study in the context of Meghalaya. In: Issues on health and healthcare in India, India studies in business and economics. Singapore: Springer; 2018. p. 379–393. Available from: http://link.springer. com/https://doi.org/10.1007/978-981-10-6104-2\_21
- Soares Magalhães RJ, Salamat MS, Leonardo L, Gray DJ, Carabin H, Halton K, et al. Mapping the risk of soil-transmitted helminthic infections in the Philippines. Knopp S, editor. PLoS Negl Trop Dis. 2015;9(9):e0003915. Available from: https://dx.plos.org/https://doi.org/10.1371/journal.pntd.0003915
- DHS. The DHS Program-India: Standard DHS, 2015–16 Dataset, Demographic Health Survey. 2017. Available from: https://dhsprogram.com/data/dataset/India\_Standard-DHS\_2015.cfm?flag=1
- Fahrmeir L, Kneib T, Lang S. Penalized structured additive regression for space-time data: a Bayesian perspective. Stat Sin. 2004;14:731–61 Available from: http://www3.stat.sinica.edu.tw/statistica/oldpdf/A14n36.pdf.
- Lang S, Brezger A. Bayesian P-splines. J Comput Graph Stat 2004;13(1):183–212. Available from: http://www.tandfonline.com/doi/ abs/https://doi.org/10.1198/1061860043010.
- 22. Besag J, Kooperberg C. On conditional and intrinsic autoregression. Biometrika. 1995;82:733–46.
- Umlauf N, Klein N, Zeileis A. BAMLSS: Bayesian additive models for location, scale, and shape (and beyond). J Comput Graph Stat 2018;27(3):612–627. Available from: https://www.tandfonline.com/doi/ full/https://doi.org/10.1080/10618600.2017.1407325.
- Spiegelhalter DJ, Best NG, Carlin BP, van der Linde A. Bayesian measures of model complexity and fit. J R Stat Soc Ser B (Statistical Methodol) 2002;64(4):583–639. Available from: http://doi.wiley.com/https://doi.org/ 10.1111/1467-9868.00353
- Goswmai S, Das KK. Socio-economic and demographic determinants of childhood anemia. J Pediatr 2015;91(5):471–477. Available from: http:// dx.doi.org/https://doi.org/10.1016/j.jped.2014.09.009
- Khanal V, Karkee R, Adhikari M, Gavidia T. Moderate-to-severe anaemia among children aged 6–59 months in Nepal: an analysis from Nepal demographic and health survey, 2011. Clin Epidemiol Glob Heal 2016;4(2):57–62. Available from: http://dx.doi.org/https://doi.org/10. 1016/j.cegh.2015.07.001
- Onyeneho NG, Ozumba BC, Subramanian SV. Determinants of childhood Anemia in India. Sci Rep. 2019;9(1):16540 Available from: http://www. nature.com/articles/s41598-019-52793-3.
- 28. Dalili H, Baghersalimi A, Dalili S, Pakdaman F, Hassanzadeh Rad A, Abbasi Kakroodi M, et al. Is there any relation between duration of breastfeeding

- and anemia? Iran J Pediatr Hematol Oncol. 2015;5(4):218–26 Available from: https://pubmed.ncbi.nlm.nih.gov/26985355/.
- National Health Mission (NHM). Ministry of Health and Family Welfare, Government of India, "6 Interventions of the Anemia Mukt Bharat Programme," 2021. Available from https://anemiamuktbharat.info/home/interventions/. Accessed 21 Sep 2021

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